# Similarity in transition to adulthood of friends and school mates 

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#### Abstract

Previous research on 'linked lives' has mainly focused on similarity in life course outcomes of parents and children, between siblings or spouses, neglecting the potentially powerful impact of friends and schoolmates. This paper uses data from the National Longitudinal Study of Adolescent and Adult Health (Add Health) to investigate the similarity of life course trajectories in the transition to adulthood of a national representative sample of young women in the US. Using recent methodological innovation in sequence analysis, we first estimate the similarity in life course trajectories among friends and peers. In the second part of the paper we combine sequence analysis to causal inference to estimate the causal effect of friends' life course transitions in respondents' transition to parenthood, marriage and cohabitation. Results indicate that friend' trajectories are more similar than random school-mates but less than siblings. Although friends seem to have a direct effect on transition to adulthood, the effect is reduced once we control for previous trajectories and other confounders.


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## 1 Introduction

Life course trajectories are embedded in social contexts. However, the importance of social interaction in life course has not been coupled with a satisfactory body of empirical research at the micro level (Balbo and Barban, 2014) Because individual lives are lived interdependently, changes in one person's life patterns often lead to changes in other people's lives as well (Elder, 1985). To what extent are life course decisions, such as family formation, divorce and separation or leaving the parental home, influenced by the behaviour of significant others?

Previous research on 'linked lives' has mainly focused on similarity in life course outcomes of parents and children, between siblings or spouses, neglecting the potentially powerful impact of friends and schoolmates. Young adults tend to turn away from their parental homes but are not yet embedded in stable relationships and therefore spend significant amounts of time with friends and schoolmates. Particularly, during this life course stage, friends function as role models and reference points for own life choices. As a possible consequence of the second demographic transition, (Lesthaeghe and Van de Kaa, 1986) friends may be equally or more important than siblings in affecting life course decisions, mainly for two reasons. First, declining fertility has led to smaller families - that is, more singletons and fewer siblings. Siblings' roles have likely been replaced by close friends. Second, friends are freely chosen by an individual and voluntary relationships have gained in importance compared to ascribed family relationships.

Our previous work (Balbo and Barban, 2014) shows that a friend's childbearing increases an individual's risk of becoming a parent. The effect is short-term and curvilinear: an individual's risk of childbearing starts increasing after a friend's childbearing, reaches its peak approximately two years later, and then decreases. We extended the previous study in three ways. Firstly, we take an holistic approach in describing the entire trajectory of transition to adulthood by examining the similarity among friends and schoolmates. Secondly, we compare friendship similarity with sibling similarity. Thirdly, we analyze the causal effect of friendship events on the respondents' trajectories to adulthood by adopting a matching design that takes into account both time-fixed confounders and pre-treatment trajectories, combining sequence analysis and propensity score matching.

We use data from the National Longitudinal Study of Adolescent and Adult Health (Add-Health) we will investigate to what degree friends share similar life course trajectories and influence each other in life course decisions. We use recent methodological developments in sequence analysis for dyadic data (Raab et al., 2014; Balbo and Barban, 2014) to describe similarity in life course trajectories and to identify peer effect in life course transitions. Furthermore, we will examine whether demographic trajectories of high school friends are more similar among the lower educated, since they do not move to college and form new friendship networks. Understanding the role of social interactions in shaping individual life course decisions may help in identifying successful policy interventions.

The first part of the paper investigates if lifecourse trajectories of friends are more similar to randomnly drawn respondents from the sample or previous schoolmates. We then extend the analysis and compare friends' similarity to siblings. The second part of the paper investigates the effect of friends' life course events (transition to parenthood; first marriage and first cohabitation) on the respondents life course trajectories.

## 2 Data

We use data from the National Longitudinal Study of Adolescent Health (Add Health) in the United States,nationally representative longitudinal survey of adolescents who were in grades 7 through 12 in Wave I (1995). The Add Health cohort (born between 1976 and 1982) was followed into young adulthood with four in-home interviews (Wave I in 1995, Wave II in 1996, Wave III in 2001 to 2002, and Wave IV in 2008 to 2009), at the end of which the sample was between 26 and 33 years old. Add Health provides an opportunity to combine three different types of information: longitudinal data on respondents' socioeconomic characteristics; information on life course events and trajectories; and data on friendship networks.

Friendship information have been collected as follows. In Wave I, in-school network information was collected and up to 10 friendship ties for each respondent were identified. In Wave III, a follow-up of the Wave I network module was administered to 3,572 respondents, who were in 7 th or 8 th grade at Wave I. In the friends module of Wave III, respondents were asked a battery of questions about their current relationship with 10 former schoolmates. These 10 people were selected into a respondent's questionnaire by a name generator based on the probability of remaining friends with that respondent. Every schoolmate selected was also part of the study sample. Following our previous work (Balbo and Barban, 2014), we used information on friendship status at Wave III to defined two typologies of network relationship: peers (i.e., former schoolmates who have never been friends) and friends (i.e., former schoolmates who became friends during high school and remained so over time).

The friend module of Add Health contains information on 27,803 dyads, 4,611 of which are listed as current friends in Wave III, 3,859 are former friends and 9,220 are pairs of individuals that are listed by the computer name generator and are not friends. Once merged with the life course trajectories of transition
to adulthood, the sample reduced to 3,100 female respondents.
[data and numbers]

## 3 Friends' similarity

### 3.1 Sequence definition

Life course trajectories are represented by monthly combinations of union and childbearing states from age 15 until wave IV (age 28-32). We designed the state space to take six possible values: Single (S); Single Parent (SP); Cohabiting (C); Cohabiting Parent (CP); Married (M) and Married Parent (MP) (Barban, 2013). The months in which family events take place are defined using retrospective questions. Sequences are calculated only for female respondents for two main reasons: first, women experience family life course events earlier than men; second, women are more reliable on the timing of life course events (e.g. pregnancy) in the Add Health sample.

We merged the life course trajectories with the friends module included in the Wave III of Add Health. We obtained information on 27,803 dyads that include friendship relationship, former friendship (respondents who were friends in the past, but discontinued their relationships) and peers (respondents nominated by the name-generator in Wave III).

Friends share more similar characteristics than peers such as geographical proximity, same racial background and family characteristics than peers (see Table 1).

| Dyad Type | Close Friends | Former Friends | Peers |
| :---: | :---: | :---: | :---: |
| Avg. Distance in miles at WI | 7.28 | 7.29 | 7.99 |
| Prop. same race | 0.82 | 0.78 | 0.71 |
| Prop. living in same state at WIII | 0.79 | 0.74 | 0.77 |
| Prop living in same county at WIII | 0.52 | 0.45 | 0.51 |
| Prop. living in same census tract at WIII <br> Prop. living in same block at WIII | 0.16 | 0.11 | 0.09 |
|  | 0.08 | 0.05 | 0.03 |
| N | 1,254 |  |  |

Table 1: Descriptives

|  |  |  |  |
| :---: | :---: | :---: | :---: |
| Dyad Type | Close Friends | Peers | Random Assigned |
| Avg. OM distance | 0.801 | 0.898 | 0.945 |
| SD OM distance | 0.470 | 0.464 | 0.455 |
| $95 \%$ CI | $(0.766-0.836)$ | $(0.874-0.920)$ | $(0.923-0.967)$ |
| N | 708 | 1,538 | 1,706 |

Table 2: Average OM sequence dissimilarity across dyadic type

|  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Dyad Type | Twins | Full-Siblings | Half Siblings | Friends |
| Avg. OM distance | 0.721 | 0.755 | 0.920 | 0.801 |
| SD OM distance | 0.477 | 0.42 | 0.397 | 0.470 |
| $95 \%$ CI | $(0.685-0.754)$ | $(0.732-0.778)$ | $(0.891-0.949)$ | $(0.766-0.836)$ |
| N | 741 | 1,251 | 442 | 708 |

Table 3: Average OM sequence dissimilarity across dyadic type

### 3.2 Sequence similarity

We use Optimal Matching Sequence Analysis to derive a measure of life course similarity across all respondents. We obtain a dissimilarity matrix, i.e. a matrix of dimension $N \times N$ containing all pairwise distances calculated with the sequence analysis algorithm.

Following the approach of (Raab et al., 2014), we use the information from the dissimilarity matrix to compute the sequence similarity for each respondent in the friend module. To avoid multiple sampling from the same person (respondents have different number of friends and peers), we randomly selected one listed friend, one peer (nominated from the computer generator) and one random respondent from the sample. We then use this information to compare how similar are life course of the selected pairs.

## 4 Friends influences in life course events

### 4.1 Matching life course sequences

In the second part of the paper, we attempt to assess the causal effect of friends' events. In particular, we analyze the effect of having a high school friends experiencing childbearing, marriage and cohabitation for the first time. Respondents have multiple friends and their effect could be cumulative. For this preliminary analysis we restrict our interest to the first friend who experienced the event.


Figure 1: Optimal Matching Distance distribution across dyadic types

In other words, respondents become at risk of childbearing (marriage; cohabitation) once a listed friend experience the event. Using the terminology common in the literature of causal inference, a respondent who has a friends experiencing the event before her, is called a treated individual. On the other hand, respondents who do not have friends who experienced the events or those who experienced the events before their friends are called controls. Treatment time depends therefore on both respondents' and friends' life course trajectories.

Our analytical approach can be described as follows:

1. Starting from dyadic data ( 27,803 friendship pairs), we identify for each respondent $i$ the earliest time when a friend $j$ experienced childbearing (marriage, cohabitation). For example, if a respondent has two friends who had a child when she was 19 and 21 , she become at risk of "friend influence" starting at age 19. In other words, her treatment time is age 19.
2. We identify all the controls, i.e. all respondents that cannot be influenced by friends because they had a child (marriage, cohabitation) before their friends
3. We calculate the propensity score of being a treated individual by estimating a probit regression.
4. We identify the pre-treatment sequence, that is the sub-sequence representing the life-course trajectory antecedent treatment time.
5. For each treated, we compare the pre-treatment sequence with all the controls by calculating Optimal Matching sequence dissimilarities.
6. Similarly, we compare the propensity scores by calculating the distances in propensity scores.
7. For each treated, we select the control who had the minimum combined distance (OM dissimilarity + propensity score)
8. Finally, we compare the proportion of individuals with children (married, cohabiting) at time $t+1, t+2, \ldots, t+10$.

### 4.2 Sequence analysis Matching

We adopt the approach described in Barban et al. (2017) to match respondents who share similar pre-treatment sequence. We use Optimal Matching dissimilar-
ity matrix with transition costs derived by the inverse of transition probabilities, i.e. transitions that are less frequent have higher costs in the Sequence analysis algorithm.

### 4.2.1 Example

Respondent $i$ is 23 when the first high school friend (listed in the friend module) has a child. She has no children and she started a cohabitation when she was 20 . At age 22 she married. We know from extensive literature on family demography that her "risk" of becoming a mother is sensibly higher than a single individual. Therefore, it would be misleading to compare this respondent with someone who is single. Ideally we want to "search" in the data if there is any other respondent who had the identical (or the most similar) life-course. In addition, we know that $i$ was born in 1982, she is Hispanic White, her family had income above the median, her parents are High School educated and so on. By combining sequence matching to propensity score, we make sure that the comparison between treated and controls takes into account all these characteristics.

### 4.3 Propensity score Matching

We base our propensity score matching on the following variables, measured before year $t$ corresponding to the year of retirement for the treatment group:

1. Year of Birth
2. Household income at Wave I (Dummy variable: 1 above median; 0 below median income)
3. Parental education (less than high school; HS or equivalent;some college;college or more; unknown)
4. Family type at Wave I (Dummy variable: 1 Living with both biological parents; 0 otherwise)
5. Race/Ethnicity (White Hispanic; Black; White Asian; White Non-Hispanic)
6. Born in the US (Dummy variable)
7. State fixed effect

Propensity scores and sequence analysis distances are normalized and than summed to obtain a unique measure of similarity among controls and treated respondents. In this particular case, we assigned equal weights to the two methods and we selected only one control as matched respondent. However as discussed in

## 5 Preliminary Results

Once obtained the matched controls, we investigate the difference in proportion of parents after the occurrence of a friend's childbearing. Figure 2 shows the difference in the proportion of parents after the friend's event before and after matching. Panel A clearly show a higher proportion of respondents who became mothers after their friends (treated) compared to the rest of the sample (controls). The difference is highly significant and would induce to assign a direct "friend" effect on the basis of these results. However, once we apply our matching procedure based on pre-treatment sequence and time-fixed propensity score, the "friend" effect seems to disappear completely. Figure 3 and 4 reports the same analysis for marriage and cohabitation. The results is similar for all the outcome with the possible exception of cohabitation, where a "negative friend effect" seems to be present in the years immediately after friends' cohabitation.

## 6 Discussion

Friends play an important role in the transition to adulthood. Friends have similar life course transitions, as they progress trough adulthood. Understanding the degree of interconnection of life course decisions such as entering a union or having a child, is important to understand how policies can be implemented during this age. Our work shed a light on the similarity of life-course transition using a unique representativ dataset on US women as they progress from adolescence into young adulthood. As noted elsewhere, this is a period demographically very dense where many life-changing events happen. The timing, the order of these events has strong links with socio-economic opportunities, later well-being and health in later life. Keenan and Grundy (2018)

The first part of the paper is dedicated to analyze if there is significant similarity among friends and other young adults. Friends are important relationships that are chosen by individuals. Therefore, friends often share similar


Figure 2: Proportion of Mothers after friends' childbearing


Figure 3: Proportion of married women after friends' marriage


Figure 4: Proportion of cohabiting women after friends' cohabitation
background and similarity, i.e. homogamy and are affected by the similar influences such as same school, neighbourhood etc. We use a distinctive feature of the Add-Health questionnaire to distinguish between friends that are in contact and may exert an influence on life course decisions and former school mates (peers) that have been exposed to similar background but may not have a direct "friend" effect on life course. We use the methodology presented by (Barban et al., 2017), and in particular sequence analysis to represent and measure life course trajectories in a holistic manner. Rather than focusing on single life course events, we take the entire trajectory as a whole. The analysis shows that friends are more similar than peers in their life trajectories. The timing, the occurrence of life course events is more similar than peers or than random individuals.

We then, extend the analysis to a sub-sample of siblings and step-siblings in the survey. We find that siblings have more in common than friends and generally have higher life course similarity. However, this is not true for stepsiblings, as they lay somewhere between friends and peers (school-mates).

In the second part of the paper, we attempt at measuring a direct effect of
friends in three life course events (childbearing, marriage and cohabitation). A first analysis, shows that when a friend experience first one of these life-changing event, her friends have higher probability of experiencing as well. However, life course events are interrelate and often the results of multiple decisions that occurred in the past. To control for each individual trajectory, we adopted a recent method that combines the literature in sequence analysis and causal inference to identify the possible causal effect of life course events. We designed a matching framework where individuals are matched on time-fixed covariates (propensity score matching) and life course trajectories (sequence analysis matching). Our preliminary results show that all the "friend" effect is masked by similar life course trajectories and other characteristics.

We aim to extend this work by looking more in dept at the variety of life course events (e.g. single parenthood, compared to parenthood in a stable union) and by stratifying the results by event at different age (i.e. teenage pregnancy). In addition, the current analysis does not take into consideration socio-economic stratification. Friends from high school may be more relevant to individuals who did not go to college and did not move far from their parents. We aim to extend the analysis by including appropriate measures of socio-economic background to identify the possible differential effect of friendship across educational groups.

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