

## **Genes related to education and frailty among older adults in the United States**

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## **Abstract**

Using data from a sample of 5,344 non-Hispanic and white adults from the Health and Retirement Study (HRS), we expand on research that links education and frailty among older adults by considering the role of genes associated with education. We calculate a genome-wide polygenic score (PGS) for education and demonstrate a strong and negative association between genes associated with education and symptoms of frailty in later life using two different indicators of frailty (Deficit accumulation and the Paulson-Lichtenberg frailty index). We also show that this association exists above and beyond years of completed education and we demonstrate that this association becomes weaker as older adults approach their 80s. Our results contribute to the education-health literature and suggest new and important pathways through which years of education may be linked to successful aging.

## **Introduction**

One of the most important areas of research for demographers involves the characterization of health and the determinants of health of aging populations (Christensen et al. 2009). One particular area that has grown over the past 25 years is research that consistently demonstrates a robust relationship between education and health (Link and Phelan 1995). The protective effect of increasing years of education is evident across a number of morbidities and is fairly consistent across different sociodemographic groups in the US (Zajacova and Lawrence 2018). In this paper, we focus on a critical aging-related phenotype, *frailty*, and evaluate the extent to which this indicator of overall health among the elderly is linked to years of education. Previous work has shown educational gradients in frailty at both the population (Etman et al, 2012, Theou et al., 2013) and individual levels (Crimmins et al, 2010, Szanton et al, 2010) but no existing work has evaluated the hypothesis that some of this association is due to genetic influences that affect education and subsequently frailty.

The purpose of the study is to explore the possibility that genetic loci that are associated with educational attainment are also implicated in the reduced likelihood of frailty in later life (60+ years). Importantly, we evaluate the possibility that genes related to education also predict frailty above and beyond the effect of educational attainment. While there is preliminary evidence that the  $PGS_{educ}$  has lasting benefits through the life course, we seek to explore how these effects manifest over biological age. That is, is the potential effect of the  $PGS_{educ}$  most predictive of frailty more or less important as individual ages? Such answers will help to elucidate the relationship effects of education on frailty by exploring the genetic mechanisms underpinning both measures.

## **Genetics and Years of Completed Education**

A recent meta-analysis of twin and family studies show that up to 40% of variance in educational attainment can be explained by genetic factors (Brainigan, McCallum & Freese, 2013).

Using genetic markers, it is now possible to quantify an individual's genetic propensity for high levels of educational attainment. One increasingly popular method is the polygenic score (PGS) approach (Dudbridge 2013), which has led to a number of advances over the past two decades in predicting disease (Jostins & Barret, 2011). Such successes are reflective of a growing consensus that complex phenotypes, such as educational attainment, are influenced by many genetic loci with very small effect sizes (i.e., highly polygenic, Visscher et al., 2017). The PGS for educational attainment (PGS<sub>educ</sub>) is a single score representing genome wide influence on academic success, as measured by formal years of schooling. Initially, social scientists developed this score (Reitvelt et al., 2013, Okbay et al., 2016, Lee et al., 2018) to explain variation in schooling due to genetic factors and the current score explains roughly 10% of the variance in education (Lee et al. 2018). In early and midlife, the PGS<sub>educ</sub> is predictive of other measures of academic success and cognitive performance (e.g., general cognitive ability, standardized exam scores, highest math course completed; Lee et al., 2018, Belsky et al., 2018). Into adulthood, the PGS<sub>educ</sub> predicts social mobility and indicators of economic success (e.g., occupational status, asset accumulation, financial stability, and wealth at retirement; Belsky et al., 2016, Belsky et al., 2018, Barth, Papageorge, and Thom, 2018).

While there is evidence that genes associated with educational attainment predict indicators of success, the relationship with indicators of *health* is less established but is growing. That is, do *genes* associated with educational attainment also predict health outcomes? Results from large consortia suggest significant genetic overlap between genes for educational attainment and a range of health indicators (e.g., Alzheimer's disease, ischemic stroke, psychiatric conditions, vascular-metabolic diseases, smoking status, and other physiological measures such as body-mass index; Bulik-Sullivan et al., 2015, Hagenaars et al., 2016, Gandal et al., 2018, Anttila et al., 2018). Indirect evidence for this association comes from two additional studies. First, Marioni et al. (2016) found that a child's PGS<sub>educ</sub> was predictive of mortality, as measured by the paternal and maternal lifespan

(Marioni et al., 2016). Second, Boardman et al. (2015) provide evidence for genetic correlation ( $r_G$ ) for genes linked to years of education and genes linked to self-rated health.

### **Aging, genes, and frailty**

From a medical perspective, frailty is a state of reduced physiological resilience and increased vulnerability to adverse events. Patients who are frail are more likely to be hospitalized, lose daily independence, have negative outcomes with medical procedures (e.g., poor response to surgeries), and have increased risk of death (Mitnitski, Mogilner, and Rockwood. 2001; Fried et al., 2001; Hamaker et al., 2012; Jung et al., 2014; Mitnitski et al. 2015; Rodriguez-Manas and Fried 2015; Hanlon et al., 2018). Frailty is an indicator of general health predicts morbidity and mortality and is considered an indicator of biological age (Fried et al. 2001; Sanderson and Scherbov, 2014; Mitnitski et al., 2017). As demographic shifts in population age-structure occur worldwide, the result is a growing population at risk for frailty, which is a debilitating and financially costly age-related syndrome (Morley et al., 2013; Goldman et al., 2013). As such, it is imperative to understand the predictors and mechanisms governing who is most at risk for frailty.

Though aging is the greatest known risk factor for frailty (Rockwood and Mitnitski, 2011), not all aging individuals become frail, and individuals develop frailty at different rates (Mitnitski & Rockwood, 2016). Those with lower levels of education are far more likely than those with relatively high levels of education to evidence symptoms of frailty at an early age (Santos-Eggimann et al. 2009). There are two competing hypotheses regarding effects of education across the life course which may explain the related gradients for frailty. The *cumulative advantage* hypothesis posits that educational gradients should increase with age (O’Rand, 2006). Alternatively, the alternative, *age as a leveler* hypothesis suggests these differences are diminished later in life (Preston, Hill, & Drevensted, 1988). To date, there is evidence for both perspectives depending on the specific measure of frailty, the age composition of the study, and whether or not frailty is measured as a static indicator of

current health or a within person measure of change in indicators of frailty over time (Yang and Lee 2010).

Bringing measured genetic polymorphisms to bear may shed light on the relevance of these two age-related models. Previous research has shown that roughly 20-40% of frailty is genetically oriented (Sanders et al. 2016; Young et al. 2016). In an oversimplification of this very complex process consider the possibility that two individuals with otherwise identical states of frailty may have arrived to this state via very different paths: one because of their environmental exposures across their lifetime and the other simply because of their genetic composition. Thus, environmentally oriented frailty and genetically oriented frailty should be considered somewhat different morbidities despite their identical appearance in terms of measurement and assessment. One way to evaluate the relevance of the two age-related models discussed above is to consider changes in the effect of the PGSed with increasing ages. If the effect of PGSed on frailty is the same across ages among adults over the age of 60 then it provides some indirect support for the cumulative advantage perspective and it suggests that some of these cumulative advantage may have origins in genes linked to educational success. On the other hand, if the effect of the educational PGS on frailty decreases with increasing age then it provides some support for the age as a leveler perspective and importantly it stresses that age even levels the playing field with respect to genetic influence on the educational determinants of frailty.

## **Material and Methods**

[Table 1 about here]

### **Data**

The data comes from the Health and Retirement Study (HRS), a biannual panel study tracking physical, emotional, and economic wellbeing during the transition into older age (Juster & Suzman, 1995, Sonnega et al., 2014). These respondents were born between 1900 and 1970 with the

interquartile range (IQR) of birth years spanning from 1930 to 1950. The HRS is sponsored by the National Institute on Aging (NIA U01AG009740) and is conducted by the University of Michigan. This study focuses on a sample of 5,344 HRS respondents with genotype data collected between 2006-2008 and that had data for both frailty measures. Descriptive statistics for all variables are provided in Table 1. Because of complications due to population stratification and the use of the  $PGS_{educ}$  values across socially identified racial and ethnic groups, our analyses are limited to those who identify as non-Hispanic and white. Descriptive statistics for all variables used in the HRS are presented in Table 1.

### **$PGS_{educ}$**

Genotypes were assessed using the Illumina HumanOmni2.5 BeadChips, with coverage of over 2.4 million genetic loci (i.e., single nucleotide polymorphisms [SNPS]). Standard quality control procedures were conducted (Ware et al., 2018). The  $PGS_{educ}$  was calculated using the sample excluding participants from the Social Science Genetic Association Consortium (SSGAC; Okbay et al., 2016). Scores were standardized within the sample.

### **Frailty indices**

Two metrics of frailty were used to assess five waves of the HRS (2004-2012). These waves were selected due to consistency in item measurement, and have also been used in previous research on frailty (Mezuk et al., 2016). When possible, items were pulled from a harmonized longitudinal file prepared by the RAND Center on the Study of Aging (RAND, 2014). When unavailable in the RAND file, matching items were identified in the biannual files. For both metrics of frailty, scores are treated as both continuous and categorical. Continuous measures reflect an unweighted count of the number of health problems (i.e., symptoms, signs, functional impairments, or abnormal

laboratory values), which are collectively referred to as “deficits.” Counting deficits stratifies respondents based on their level of functional decline, and thus, their degrees of vulnerability (Rockwood et al., 2005). Categorical measures were treated as binary and are based on clinical cutoff that have emerged to classify an individual’s *frailty status* (i.e., frail or non-frail). Importantly, frail individuals have increased mortality compared to individuals of the same age (Vaulpel, Manton, & Stallard, 1979). While frailty is considered a reversible or dynamic state transitioning back from frailty is not common (Gill et al. 2006). The explanation for this is that once enough homeostatic reserve is lost it cannot be regained.

### **Paulson-Lichtenberg frailty index (PLFI)**

As described by Paulson and Lichtenberg (2015), the frailty phenotype was designed to be a best-fit representation of the Fried et al. 2001 frailty phenotype using HRS data. The PLFI included five symptoms: wasting (i.e., individual reported loss of at least 10% of body weight over a 2-year period), weakness (i.e., “Because of health problems, do you have any difficulty with lifting or carrying weights over 10 pounds, like a heavy bag of groceries”), slowness (i.e., “Because of a health problem, do you have any difficulty with getting up from a chair after sitting for long periods), fatigue (i.e., “Have you had any of the following persistent or troublesome problems: sever fatigue or exhaustion?”), and falls (“Have you fallen down in the past 2 years”). The continuous PLFI measure had a range of 0-5, whereas the binary measure of frailty status used the established cut off of PLFI >3 as indicative of “frail” (Paulson & Lichtenberg, 2015).

### **Deficit accumulation model frailty index (FI)**

Frailty in the HRS was also operationalized using the deficit accumulation model put forth by Rockwell et al. 2005, and Mitniski, Song, and Rockwood, 2013. The FI is a calculated ratio of health



deficits out of the total number of possible deficits (e.g., 10 deficits present / 30 possible deficits indicates a FI of 0.33). The FI developed for use in the HRS uses 30 self-report health measures and was recreated using the item list published in Mezuk et al., 2016. Items included a variety of deficits such as difficulties with activities of daily living (e.g., dressing, bathing, toileting, cooking, shopping, changes activities of daily living), problems with pain (e.g., general pain, back pain, headache), problems with worsening memory or dementia, disturbances in sleep (e.g., persistent fatigue, problems falling or staying asleep), motor impairment (e.g., falling, impaired mobility, gross-motor impairment, fine motor impairment), and presence of health conditions (e.g., incontinence, cancer, arthritis, psychiatric conditions, depression, lung disease, respiratory problems, diabetes, stroke, angina, heart failure, heart attack, or high blood pressure). The continuous FI is a proportion ranging from 0-1, while the established cutoff of  $FI > 0.25$  defines frailty status (Rockwell et al. 2005).

### **Statistical analysis**

Statistical analyses were conducted using R (version 3.4.1, Project for Statistical Computing; R Core Team, 2015) and Stata (StataCorp, 2011). Simple association of frailty score with polygenic score was done by chi-squared test of independence. Because frailty was repeatedly assessed we use all data across waves. In total we had 827 individuals with one observation, 1186 with two, 3330 with three, and 1 person had four frailty observations. Multi-level regression with the *xtmixed* procedure in Stata was used to model the association of  $PGS_{educ}$  and frailty score in which observations are nested within people. We report the intra class correlation coefficient (ICC) for each model. Similarly, multi-level logistic regression was used with the *xtmelogit* command in Stata to model  $PGS_{educ}$  and frailty status (frail / non-frail). Level 1 error variance was assumed to be  $\frac{\pi^2}{3}$  (Guo and Zhao 2000). All models control for the top five principle components (PCs) to control for effects of population

stratification or spurious association due non-causal allele frequency differences across ancestry groups (Patterson, Price, & Reich, 2006).

## Results

### [Table 2 about here]

Table 2 presents bivariate associations between  $\text{PGS}_{\text{educ}}$  (cut into quartiles) and the four metrics of frailty. These results provide the first evidence for our hypothesis linking genes related to educational outcomes and the indicators of frailty among older adults. The PLFI is the highest among those in the 1<sup>st</sup> quartile (lowest level) of the education  $\text{PGS}_{\text{educ}}$  ( $\bar{x} = 1.29$ ) and the least frail are those in the 4<sup>th</sup> quartile of the PLFI ( $\bar{x} = 1.18$ ). This difference is a small effect size ( $d = .09$ ) but it remains statistically and substantively significant (the p-value for this test is derived from a one-way ANOVA test and demonstrates that this relationship is statistically significant ( $p < .001$ ). We show comparable estimates using the FI, a more comprehensive measure of frailty in which again, those with the lowest education  $\text{PGS}_{\text{educ}}$  have a frailty score ( $\bar{x} = 24.41$ ) that is significantly higher than the frailty score among those with the highest education  $\text{PGS}_{\text{educ}}$  ( $\bar{x} = 23.07$ ). As with the PLFI, the effect size remains quite small ( $d = .11$ ) but important nevertheless. Table 2 also evaluates the same associations but at their respective thresholds to identify a state of frailty. While only 14.68% of those with the highest education  $\text{PGS}_{\text{educ}}$  were in a frail state according to the PL score 17.08% of those with the lowest education  $\text{PGS}_{\text{educ}}$  scores were in a frail PLFI state and this same association is seen with the threshold for the frailty index. Specifically, while 40.27% of those with low education  $\text{PGS}_{\text{educ}}$  were frail at the time of the interview, only 36.72% of those with high education  $\text{PGS}_{\text{educ}}$  scores were frail. As with the continuous indicators, both binary assessments of frailty status were statistically significant ( $p < .025$  and  $< .008$ , respectively).

### [Table 3 about here]

Table 3 presents the results from multilevel analyses for the two continuous indicators of frailty described above in which observations are nested within individuals. For each measure we present the results of three models. The first regresses the frailty score on the education PGS<sub>educ</sub> with controls for gender, age, and the top 5 PCs. The second introduces a control for years of education and the third introduces the interaction between the PGS<sub>educ</sub> and age to evaluate the possibility of age related changes in the influence of the PGS<sub>educ</sub> on frailty. Model 1 provides a baseline indicator for the effect of the education PGS<sub>educ</sub> on PLFI. The effect ( $b = -.06$ ) suggests that a one standard deviation increase in the education PGS<sub>educ</sub> reduces frailty by .06 points. As with the bivariate associations this denotes a small effect size ( $d = .05$ ) but again it remains a statistically significant and meaningful association. Model 2 introduces years of completed education as a mediating mechanism to link education PGS<sub>educ</sub> to frailty. As expected, the PGS<sub>educ</sub> coefficient is significantly reduced ( $b = -.03$ ) but most importantly, the association remains statistically significant. This provides evidence for our primary hypothesis. Namely, that genes linked to educational attainment provide health protections that are *above and beyond* each additional year of formal education. Model 3 evaluates our interest in this association as individual's age. Accordingly, we introduce an interaction between age and the education PGS<sub>educ</sub>. As shown in Model 3, the interaction is positive and significant which suggests that the protective effects of the education PGS<sub>educ</sub> on frailty diminish with age. To better gauge the meaning of this interaction, we used the postestimation *margins* command in Stata to retrieve parameter estimates and confidence intervals for age specific slopes for the education PGS<sub>educ</sub> - Frailty association. These estimates are shown graphically in Figure 1a. The horizontal line represents the null hypothesis that there is no association between education PGS<sub>educ</sub> and frailty. We show the largest effects ( $b = -.06$ ) among those who are 65 years old, the average effect ( $b = -.03$ ) among those who are 75, and non-significant associations among those who are 80 and 85.

**[Figure 1 about here]**

A very similar story is shown in Models 4-6 in which the dependent variable is now the deficits accumulation model or FI. Here, the baseline association ( $b = -.84$ ) is evident above and beyond controls for age, gender, and the five principal components and this effect is nearly cut in half after controlling for years of education in Model 5 ( $b = -.44$ ). As with the PLFI, this association remains statistically significant despite controlling for education. Similarly, the association is reduced as people age as indicated by the positive interaction in Model 6 ( $b = .24$ ) which is statistically significant. The functional form of this interaction and the age thresholds at which education  $PGS_{educ}$  is linked to frailty is nearly identical to the PL score despite using different and more comprehensive indicators of frailty.

The results presented in Table 4 replicate the models in Table 3 but use a binary indicator of a frailty status with an a priori threshold for each measure described above. As with the continuous indices, the education  $PGS_{educ}$  is significantly associated with reduced likelihood of frailty using the established cutoffs for the PLFI ( $b = -.16$ ) and the FI ( $b = -.21$ ). Statistical controls for years of completed education reduce these associations by more than one-half for both PLFI ( $b = -.07$ ) and FI ( $b = -.09$ ) and rendering the main effect for the PLFI to drop below statistical significance (95% CI  $[-.16, .03]$ ). As with the continuous indicators, the interaction with age is both positive and statistically significant. Figures 1c and 1d plot the fitted probability of each respective indicator of frail status as a function of age and education  $PGS_{educ}$  in which the solid line is among those with the lowest education  $PGS_{educ}$  (-2SD) and the dashed line is among those with the highest education  $PGS_{educ}$  (+2SD). In both cases, the story is nearly identical to the continuous indicators in which the effect is the most evident among those ages 75 and younger and non-existent among those 80 and older.

## **Discussion**

A breadth of work has established a link between educational attainment and health (Zajakova and Lawrence 2018). Recently, attention has been directed at extending this research to study the link between educational attainment and an age-related health outcome, frailty. To the best of our knowledge, this is the first study to explore whether *genes* related to educational attainment directly or indirectly influence health and aging, as measured by frailty. Specifically, we hypothesized that the PGS<sub>educ</sub> would be predictive of frailty, and this association would exist above and beyond the direct effects of actual educational attainment. That is, if the genetic mechanisms are dependent on education then we would expect that the association of the PGSscore would be rendered non-significant when controlling for education in the model. Alternatively, if the genetic mechanisms are influencing health above and beyond years of education then we should continue to see a significant association between PGS<sub>educ</sub>. Indeed, that is what we observed. Across two frailty metrics we found that genes related to educational attainment predict frailty. That is, participants with higher PGS<sub>educ</sub> scores had better general health (i.e., less functional decline) compared to participants with lower scores. This held true for both a comprehensive frailty index (FI) and the measure reflecting the frailty phenotype described by Fried (2002; PLFI). Similarly, participants with higher PGS<sub>educ</sub> were less likely to be classified as “frail” based on established cut-offs. Standing alone, this finding may not be entirely surprising. As the PGS<sub>educ</sub> is predictive of actual educational attainment, it is possible that the genetic effects operate indirectly through exposure to education. Thus, we explored an alternative explanation in which the PGS<sub>educ</sub> operates independently of actual educational attainment.

While we did not explore other intermediate phenotypes between the PGS<sub>educ</sub> and frailty, several biological, cognitive, social, or psychosocial mediators have been put forth linking education to health. The bulk of these involve benefits derived from years of completed education. Accordingly, the residual effect of PGS<sub>educ</sub> and frailty suggests that there is something critical in the

education-health process that also operates above and beyond years of completed education. In their comprehensive review Zajakova and Lawrence (2018: 275) discuss what they call a “signaling or credentialing perspective” in which the attainment of specific degrees provides new sources of human and social capital that frames an individual as productive or skilled. Accordingly, it may be that individuals with higher PGSed scores present themselves in a manner that is concordant with a successful or productive individual and thus create a response from others in which these signals lead to a positive framing of an individual that exists above and beyond their years of education. This is what is referred to as an evocative gene-environment correlation because an individual’s cumulative genotype may evoke environments that are associated with both increasing education and reduced risks of frailty. This perspective is in line with our findings but our analyses cannot rule out other explanations.

Other explanations include but are not limited to the following. The education PGS could be predicting health which then influences years of education. There is a large and growing body of work characterizing the rate of biological aging differences across social groups with a focus on education (Sanderson and Scherbov 2014; Mintinski, Howelett, and Rockwood 2017) and it is possible that those with higher education PGS scores are simply aging at a slower rate than those with lower PGS scores. A critical psychological mechanism could be the role of self-efficacy or mastery in which individuals with higher educational PGS may also have higher levels of efficacy that lead to both educational success, increased resilience, and delayed frailty onset (Stretton et al. 2006). This same perspective is supported by work linking education to hopelessness and other indicators of sense of control (Mitchell et al. 2016). It is possible that the education PGS is involved in complex biological and psychiatric processes that reduce the sense of hopelessness and subsequently increase the likelihood of completing education. This same reduced hopelessness is then instrumentally linked to reduced onset of frailty. We encourage future researchers to consider these

pleiotropic mechanisms but also the evocative rGE explanation in future work to better understand the cognitive, social, and psychological pathways through which genes associated with education may reduce the likelihood of frailty regardless of one's level of education.

Another key finding of our study was that the association between  $PGS_{educ}$  scores and frailty diminished with age. That is to say, genes related to educational attainment are more strongly predictive of frailty for participants in their 60's and 70's, and the effect was nearly absent beyond age 80. This is an important finding that is consistent with previous analyses of age trends in allostatic load (Crimmins et al., 2003), where age-related increases in allostatic load flattened at the oldest ages (Seeman et al. 2008). For education-related differences in physical performance tasks (e.g., grip strength, balance, walking speed, and chair stands), the effect also diminishes at advanced ages (80+ years). Given that the direct effect of education on health diminishes with age, we would expect that genes associated with education would also have less of an effect on health at advanced ages. Several explanations, ranging from sociological to biological, supporting this expectation have been put forth. First, differences in increasing risk of frailty with age have been documented across birth cohorts within the HRS (Yang and Lee, 2010). However, when birth cohorts were combined Yang and Lee (2010) found evidence that rates of frailty decrease with age. Secondly, phenotypic variation increases with age, thus any predictors of late-life/aging phenotypes weaken with progressing age. Rates of frailty were found to decrease with age, and it is well known that mortality rates slow at advanced ages (80+ years, Greenwood and Irwin 1939; Horiuchi and Wilmoth 1998; Thatcher 1999; Thatcher, Kannisto and Vaupel 1998; Gavrilova et al., 2017). Thirdly, this may reflect a survivor effect wherein those with worse physical function and lower education die at younger ages and are not included in the sample of aged participants (Welmer et al., 2013). Lastly, the relative importance of education on physical performance is less at older ages due to accumulation of deficits. Overall, the result that  $PGS_{educ}$  is less predictive of frailty beyond age 80

when birth cohorts are combined is supportive of the *age as leveler* hypothesis; although, attention should be paid to intracohort heterogeneity (Yang and Lee, 2010, Vaupel, Manton, Stallard, 1979).

There are several strengths of our study. While frailty is an easy concept to define, it is surprisingly difficult to measure and therefore, difficult to study. There is no consensus on the optimal way to measure frailty (Cigolle et al., 2009; Chao et al., 2018), however, there are two predominant clinical frailty metrics, the FI and the PLFI. The utility of these frailty metrics has been validated in different populations (Yang and Lee, 2010, Wu et al., 2018). While separate indices are not always convergent, they are often treated as complementary as they all are valid predictors of subsequent morbidity and mortality (Cesari et al., 2013, Theou et al., 2013b, Blodgett et al., 2015, Theou et al., 2015). A major strength of this study is that we use two frailty metrics and find converging evidence for the independent effect of the  $PGS_{educ}$ , above and beyond years of education, on frailty. Similarly, we find that both models of frailty support the *age as a leveler* hypothesis for the effects of genes associated with education across the life course. Finally, few studies have attempted to parse the effects of education from underlying genetic propensity for educational attainment. Unlike a majority of work, our study begins to disentangle this effect to illuminate possible underlying mechanisms.

Our study contributes to the larger body of work for studying the educational effects on aging, however, some limitations should be considered. We did not assess cohort differences in frailty dynamics. Others have shown that the genetic associations vary across birth cohort (Domingue et al. 2016) and historical periods (Boardman et al. 2010). Thus, the predictive power of a PGS in a replication sample may be attenuated with differences across studies in genotyping platforms, environmental contexts, or demographic characteristics (e.g., age, cohort, sex, or ethnicity, Wray et al., 2013, Ware et al., 2017, Tropf et al., 2017, Boardman et al., 2018). That is, the  $PGS_{educ}$  score may be most predictive in replication samples that closely match the discovery sample.



As the PGS was optimized for prediction in populations of European ancestry and may not replicated in other samples (Carlson et al., 2013, Martin et al., 2017), we too limited our sample. We encourage future work to evaluate comparable hypotheses with the full sample of respondents from the HRS.

Finally, we evaluated our hypotheses with men and women together. Others have shown gender differences in genetic associations with factors associated with frailty such as psychological resilience (Boardman et al. 2008). We simple controlled for gender in our models to but we did not specifically evaluate the possibility that these associations are systematically different as a function of gender identity. Compared to men, women tend to evidence higher levels of frailty, become frail earlier in the life course, but the effect of frailty on mortality is significantly less among women compared to men suggesting the composition of frail individuals may be very different for older men and women (Germain et al. 2016; Gordon et al. 2017) We encourage future work to consider the role of gender as a key mechanism linking education PGS, education, frailty, and survival.

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**Table 1.** Descriptive statistics for all variables used in the analyses (n = 5,344)

	Mean/%	SD/N
Frailty Measures		
PL Score	1.23	1.18
PL Score > 3	0.15	828
Frailty Index	23.59	12.66
Frailty Index >=25	37.85	2023
Female	0.58	3106
Age (years)	74.66	6.86
Education (years)	13.04	2.55
Education PGS (z)	0.00	1.00
PC1 (z)	0.00	1.00
PC2 (z)	0.00	1.00
PC3 (z)	0.00	1.00
PC4 (z)	0.00	1.00
PC5 (z)	0.00	1.00

Note: Data come from the Health and Retirement Study (HRS).

**Table 2.** Bivariate associations between education PGS and frailty

	Quartile of Education PGS				pr. <
	Q1	Q2	Q3	Q4	
PL Score	1.29	1.25	1.19	1.18	0.001
PL Score > 3	17.08	15.39	14.77	14.68	0.025
Frailty Index	24.41	23.79	23.08	23.07	0.000
Frailty Index >=25	40.27	37.67	36.72	36.72	0.008

Note: Data come from the Health and Retirement Study (HRS).

PGS<sub>educ</sub>, Education, & Frailty

**Table 3.** Education related genes and continuous indicators of frailty

	PL Frailty Index						Frailty 100 Index					
	Model 1		Model 2		Model 3		Model 4		Model 5		Model 6	
Female	0.34	(0.28, 0.39)	0.31	(0.26, 0.37)	0.32	(0.26, 0.37)	2.36	(1.75, 2.97)	2.10	(1.50, 2.71)	2.11	(1.51, 2.72)
Age	0.29	(0.26, 0.31)	0.28	(0.26, 0.30)	0.28	(0.25, 0.30)	3.49	(3.27, 3.70)	3.40	(3.18, 3.61)	3.39	(3.18, 3.60)
PC1	-0.01	(-0.04, 0.01)	-0.01	(-0.04, 0.01)	-0.01	(-0.04, 0.01)	-0.04	(-0.35, 0.26)	-0.02	(-0.32, 0.28)	-0.02	(-0.32, 0.28)
PC2	0.01	(-0.02, 0.03)	0.01	(-0.01, 0.04)	0.01	(-0.01, 0.04)	0.22	(-0.08, 0.52)	0.26	(-0.04, 0.56)	0.26	(-0.04, 0.56)
PC3	0.00	(-0.03, 0.02)	-0.01	(-0.03, 0.02)	-0.01	(-0.03, 0.02)	-0.18	(-0.49, 0.12)	-0.26	(-0.56, 0.04)	-0.26	(-0.56, 0.04)
PC4	-0.01	(-0.04, 0.02)	-0.01	(-0.04, 0.02)	-0.01	(-0.04, 0.02)	-0.13	(-0.43, 0.18)	-0.15	(-0.44, 0.15)	-0.14	(-0.44, 0.15)
PC5	0.04	(0.01, 0.06)	0.04	(0.01, 0.06)	0.04	(0.01, 0.06)	0.57	(0.27, 0.87)	0.56	(0.26, 0.86)	0.56	(0.26, 0.86)
Education PGS	-0.06	(-0.09, -0.04)	-0.03	(-0.06, -0.01)	-0.03	(-0.06, 0.00)	-0.84	(-1.14, -0.54)	-0.44	(-0.75, -0.13)	-0.41	(-0.72, -0.10)
Education PGS*Age			-0.13	(-0.16, -0.10)	-0.13	(-0.16, -0.1)			-1.67	(-1.98, -1.36)	-1.67	(-1.98, -1.36)
Intercept					0.03	(0.00, 0.05)					0.24	(0.02, 0.45)
	1.05	(1.01, 1.09)	1.06	(1.02, 1.10)	1.06	(1.02, 1.10)	22.46	(21.99, 22.93)	22.63	(22.17, 23.09)	22.61	(22.15, 23.08)
level 1	0.83	(0.82, 0.84)	0.83	(0.76, 0.81)	0.83	(0.82, 0.85)	7.22	(7.11, 7.34)	7.23	(7.12, 7.34)	7.23	(7.12, 7.34)
level 2	0.79	(0.77, 0.82)	0.78	(0.76, 0.81)	0.78	(0.76, 0.81)	10.07	(9.83, 10.31)	9.93	(9.69, 10.17)	9.92	(9.69, 10.16)
Error Variance												
Level 1	0.69	(0.67, 0.71)	0.69	(0.58, 0.65)	0.69	(0.67, 0.71)	52.20	(50.58, 53.86)	52.26	(50.64, 53.92)	52.26	(50.65, 53.93)
Level 2	0.63	(0.59, 0.67)	0.61	(0.58, 0.65)	0.61	(0.58, 0.65)	101.41	(96.69, 106.37)	98.59	(93.97, 103.45)	98.45	(93.83, 103.30)
Rho	0.48		0.47		0.47		0.66		0.65		0.65	

Note: Data come from the Health and Retirement Study. Cell entries represent parameter estimates from a multilevel model in which observations are nested within individuals. Values in parentheses represent 95% confidence intervals.

PGS<sub>educ</sub>, Education, & Frailty

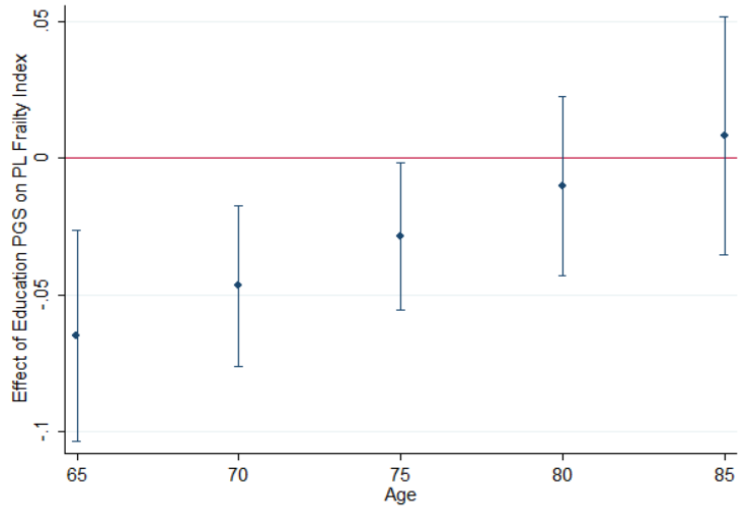
**Table 4.** Education related genes and binary indicators of frailty

	Model 1		PL Frail Model 2		Model 3		Model 4		Frail 100 Index Model 5		Model 6	
	Female	0.99	(0.79, 1.19)	0.94	(0.74, 1.14)	0.94	(0.74, 1.14)	0.70	(0.50, 0.90)	0.63	(0.43, 0.82)	0.63
Age	0.79	(0.69, 0.88)	0.76	(0.67, 0.85)	0.76	(0.67, 0.85)	0.86	(0.77, 0.94)	0.82	(0.73, 0.91)	0.82	(0.73, 0.91)
PC1	-0.02	(-0.11, 0.07)	-0.01	(-0.11, 0.08)	-0.01	(-0.11, 0.08)	0.01	(-0.09, 0.10)	0.01	(-0.08, 0.11)	0.01	(-0.09, 0.11)
PC2	0.01	(-0.08, 0.11)	0.02	(-0.07, 0.12)	0.02	(-0.07, 0.12)	0.04	(-0.06, 0.13)	0.05	(-0.05, 0.15)	0.05	(-0.05, 0.15)
PC3	-0.02	(-0.12, 0.07)	-0.04	(-0.13, 0.05)	-0.04	(-0.13, 0.05)	0.02	(-0.08, 0.12)	0.00	(-0.10, 0.10)	0.00	(-0.10, 0.10)
PC4	-0.02	(-0.12, 0.07)	-0.03	(-0.12, 0.07)	-0.03	(-0.12, 0.07)	-0.02	(-0.12, 0.07)	-0.03	(-0.12, 0.07)	-0.03	(-0.12, 0.07)
PC5	0.17	(0.08, 0.27)	0.17	(0.07, 0.27)	0.17	(0.07, 0.27)	0.19	(0.09, 0.29)	0.19	(0.09, 0.29)	0.19	(0.09, 0.29)
Education												
PGS	-0.16	(-0.26, -0.07)	-0.07	(-0.16, 0.03)	-0.08	(-0.18, 0.00)	-0.21	(-0.31, -0.11)	-0.09	(-0.19, 0)	-0.08	(-0.18, 0.00)
Education			-0.40	(-0.50, -0.30)	-0.40	(-0.50, -0.30)			-0.49	(-0.59, -0.39)	-0.49	(-0.59, -0.39)
PGS*Age					0.11	(0.03, 0.19)					0.14	(0.07, 0.22)
Intercept	-3.51	(-3.74, -3.29)	-3.49	(-3.71, -3.27)	-3.50	(-3.72, -3.28)	-1.40	(-1.56, -1.24)	-1.36	(-1.51, -1.2)	-1.37	(-1.53, -1.21)
Error Variance												
Level 2	4.77	(4.08, 5.58)	4.64	(3.97, 5.44)	4.61	(3.94, 5.4)	7.71	(6.89, 8.64)	7.47	(6.66, 8.38)	7.46	(6.65, 8.36)
Rho	0.59		0.59		0.58		0.70		0.69		0.69	

Note: Data come from the Health and Retirement Study. Cell entries represent parameter estimates from a multilevel model in which observations are nested within individuals. Values in parentheses represent 95% confidence intervals.

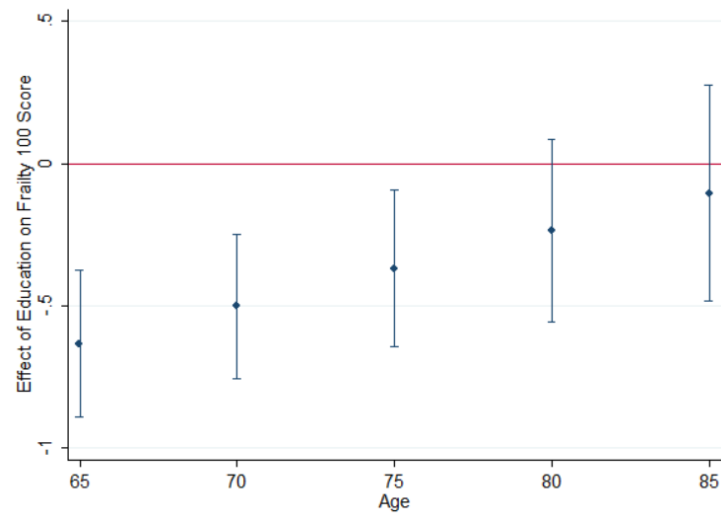
## PGS<sub>educ</sub>, Education, & Frailty

Figure 1a. PGS\*Age interaction PL Frailty



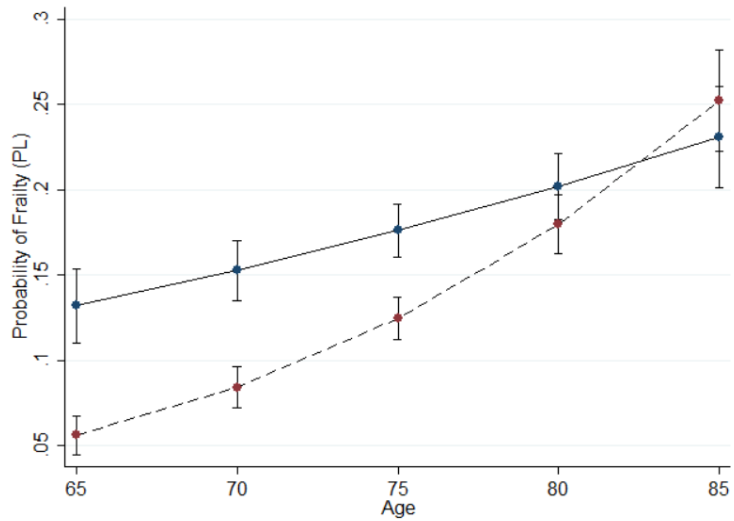
Note: Estimates derived from Table 3 Model 3.

Figure 1b. PGS\*Age interaction Frail 100 Index



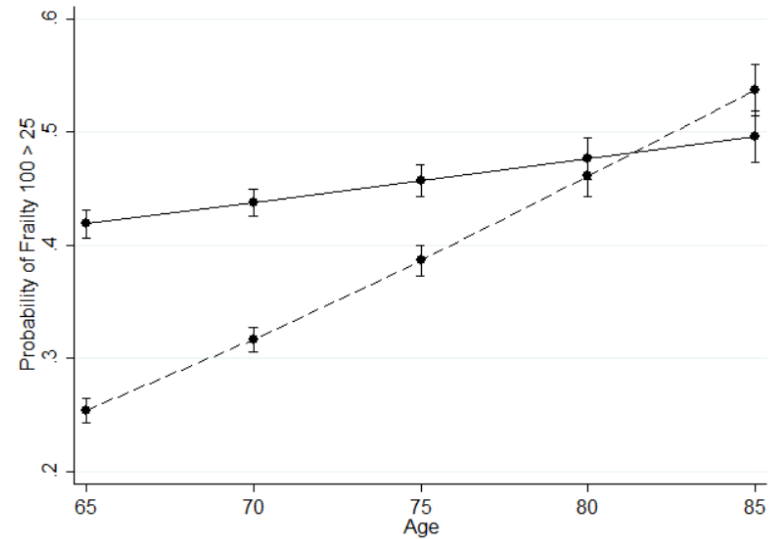
Note: Estimates derived from Table 3 Model 6.

Figure 1c. PGS\*Age interaction on risk of frail status (PL)



Note: Estimates derived from Table 4 Model 3.

Figure 1d. PGS\*Age interaction on risk of frail status (F100)



Note: Estimates derived from Table 4 Model 6.